$$-\int \frac{E\left\{D_h(X)I(T\geq t)\exp(\beta_A W(t))\frac{I(C\geq t)}{G(t|X)}\right\}}{E\left\{Y_A(t)\exp(\beta_A W(t))\right\}}dM_G(t)(3.17)$$

(Note that we can set  $\Delta/\bar{G}(T \mid X)$  to 1 in this formula again).

As remarked earlier, since C can be discrete on an arbitrarily fine grid, the projection (3.16) for the discrete multiplicative intensity model also holds for the continuous multiplicative intensity model under appropriate measurability conditions. For a formal treatment of the continuous case, see van der Vaart (2001). In Chapter 6 we state an immediate generalization (Lemma 6.1) of this lemma to multiplicative intensity models for a general counting process A(t).

Proof. The statements up to and including (3.15) were proved in Section 2.2 of Chapter 2. For the projection onto  $T_{CAR}$ , we refer to Theorem 1.1, which was only formally proved for the case where C is discrete. Thus, we only need to show the last projection result (3.17) and the general formulas  $g(H_{V,1})(t)$  and  $g(H_{V,2})(t)$ . In this proof, we will now and then suppress the index A that we used for the parameters of the censoring intensity  $E(dA(t) \mid \mathcal{F}(t))$ . Firstly, consider the special case  $V(C,X) = U_G(D_h)(Y) = D_h(X)I(C > T)/\bar{G}(T \mid X)$ . Since  $H_{V,1} = 0$ , we have

$$H_V(t,\mathcal{F}(t)) = -E(D_h(X)\Delta/\bar{G}(T\mid X)\mid C > t,\bar{X}(t)).$$

Let us denote the expectation with  $f(\bar{X}(t))$ . Thus, the numerator of  $-g(H_V)(t)$  is given by  $E(E(D_h(X)\Delta/\bar{G}(T \mid X) \mid C > t))$  $t, \bar{X}(t) Y_A(t) \exp(\alpha W(t))$ . We have  $E(f(\bar{X}(t)) Y_A(t) \exp(\alpha W(t)))$  $E(f(\bar{X}(t))I(T \geq t)\bar{G}(t \mid X) \exp(\alpha W(t)))$ , where we use  $Y_A(t) = I(C \geq t)$  $t, T \geq t$ ). We can move  $I(T \geq t)\bar{G}(t \mid X) \exp(\alpha W(t))$  inside the conditional expectation of  $f(\tilde{X}(t))$ . The resulting term can now be rewritten as follows:

$$E(E(D_h(X)\Delta \bar{G}(t\mid X)I(T\geq t)\exp(\alpha W(t))/\bar{G}(T\mid X)\mid C>t,\bar{X}(t)))$$

$$= E(E(D_h(X)I(T \geq t) \exp(\alpha W(t)) \mid C > t, \bar{X}(t)))$$

$$= E(E(D_h(X)I(T \ge t) \exp(\alpha W(t)) \mid \bar{X}(t)))$$

$$= E(D_h(X)I(T \ge t)\exp(\alpha W(t))).$$

At the first equality, we conditioned on X and C > t and used that  $E(\Delta)$ C > t, X =  $\bar{G}(T \mid X)/\bar{G}(t \mid X)$ , and at the second equality we use that, by CAR, for any function V(X)  $E(V(X) \mid C > t, \bar{X}(t)) = E(V(X) \mid \bar{X}(t))$ . Finally, for the purpose of estimation of this last expectation we note that

$$E(D_h(X)I(T \geq t) \exp(\alpha W(t))) = E(D_h(X) \exp(\alpha W(t)) \Delta/\bar{G}(T \mid X)),$$

which proves (3.17).

For a general choice of V = V(X, C) we have  $g(H_V) = g(H_{V,1}) - g(H_{V,2})$ , where  $H_{V,1} = E(V(X,C) \mid C = t, \tilde{X}(t))$  and  $H_{V,2} = E(V(X,C) \mid C > t)$  $t, \bar{X}(t)$ ). The numerator  $E(E(V(X,C) \mid C > t, \bar{X}(t))Y_A(t) \exp(\alpha \hat{W}(t)))$  of  $q(H_{V.2})$  is given by

$$E(E(V(X,C)I(T \ge t)\bar{G}(t \mid X) \exp(\alpha W(t)) \mid C > t, \bar{X}(t))),$$

Let  $Q_{V,2}(\bar{X}(t)) = E(V(X,C)I(T \ge t)\bar{G}(t \mid X) \exp(\alpha W(t)) \mid C > t, \bar{X}(t)).$ Then the last term is given by

$$E(Q_{V,2}(\bar{X}(t))I(C \geq \min(t,T))/\bar{G}(\min(t,T) \mid X)).$$

The numerator  $E(E(V(X,C)\mid C=t,\bar{X}(t))Y_A(t)\exp(\alpha W(t)))$  of  $g(H_{V,1})$ 

$$EQ_{V,2}(\bar{X}(t)) = E(E(V(X,t)\bar{G}(t\mid X)I(T\geq t)\exp(\alpha W(t))\mid C=t,\bar{X}(t))).$$

By CAR, we have for any function D(X) that  $E(D(X) \mid C = t, \bar{X}(t)) =$  $E(D(X) \mid \bar{X}(t))$ . Thus, the numerator  $E(V(X,t)\bar{G}(t \mid X)I(T) \geq$  $(t) \exp(\alpha W(t)))$  of  $g(H_{V,2})$  is actually given by

$$E(V(X,t)\bar{G}(t\mid X)I(T\geq t)\exp(\alpha W(t))I(C\geq \min(t,T))/\bar{G}(\min(t,T)\mid X)).$$

Optimal Mapping into Estimating Functions

The mapping from  $D_h$  into observed data estimating functions is a sum of two mappings  $IC_0$  and  $IC_{CAR}$ , where  $IC_0$  is an initial mapping satisfying  $E_G(IC_0(Y \mid G,G) \mid X) = D(X)$   $F_X$ -a.e. for D in a non empty subset  $\mathcal{D}(
ho_1(F_X),G)\subset \mathcal{D}$  and  $IC_{CAR}$  is the projection of  $IC_0$  onto the tangent space  $T_{CAR}(P_{F_X,G})$ . By Theorem 1.1, we have

$$T_{CAR}(P_{F_X,G}) = \overline{\left\{\int Q(u,ar{X}(u))dM_G(u):Q
ight\}} \subset L^2_0(P_{F_X,G}),$$

where

$$dM_G(u) = I(\tilde{T} \in du, \Delta = 0) - I(\tilde{T} \ge u)\Lambda_C(du \mid X)$$

is the martingale of  $A(\cdot) = I(C \le \cdot)$  w.r.t. history  $\mathcal{F}(t) =$  $\sigma(\bar{A}(t-), \bar{X}(\min(C,t-)))$ . Here  $Q(u, \bar{X}(u))$  ranges over all functions for which  $\int Q(u, \bar{X}(u))dM_G(u)$  has finite variance so that it is an element of  $L_0^2(P_{F_X,G})$ . By Theorem 1.1 (for the discrete case) and van der Vaart (2001) (for the continuous case), we also have for any  $D \in \mathcal{D}$ 

$$IC_{CAR}(Y \mid Q(F_X, G), G, D) = \int Q(F_X, G)(u, \bar{X}(u))dM_G(u),$$